# Supplementary Materials for 

# Resilience and Recovery of Overexploited Marine Populations 

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This PDF file includes:

Figs. S1 to S7
Tables S1 and S2
References

## Materials and Methods

## Definitions

We used the following operational definitions of depletion, recovery, and recovered. A stock was classified as depleted when its biomass (spawning stock or total biomass) fell below half of its MSY reference point, $\mathrm{B}_{\text {MSY }}$. This definition aligns with the legal definition of overfished in the USA (30) and Australia (31), and is meant to allow for stock fluctuations below $B_{\text {MSY }}$. We considered a stock to be recovered once its biomass exceeded $B_{M S Y}$. $B_{\text {MSY }}$ is a legislated rebuilding target in at least the USA, and has been proposed as such in Europe. Maintaining stocks at or above $\mathrm{B}_{\mathrm{MSY}}$ is also a legal obligation for signatories to the United Nations Convention on the Law of the Sea (UNCLOS). Recovery is then the demographic process of population growth between the time of depletion, $\mathrm{B}<0.5 \mathrm{~B}_{\mathrm{MSY}}$, and the time when the biomass first exceeds $\mathrm{B}_{\text {MSY }}$. Although we do not use an explicit definition of overfishing for our analysis, we can define overfishing according to the above definitions as a fishing mortality that leads to a depleted stock (theoretically, according to our definitions, $\mathrm{F}>2 \mathrm{~F}_{\mathrm{MSY}}$ ), and/or is too high to allow for recovery to $\mathrm{B}_{\text {MSY }}$ (in theory, $\mathrm{F}>\mathrm{F}_{\text {MSY }}$ ). Note that according to these definitions a stock may be classified as overfished more than once throughout its exploitation history and we term each such event a depletion event (see Fig. S3 for an illustration).

## Data

Our analyses are based upon Version 1.0 of the Ram Legacy Stock Assessment Database (11), with European stock assessments conducted by International Council for the Exploration of the Seas (ICES) updated to 2011. For each stock, we obtained time series of estimated biomass B (spawning stock or total biomass, in order of preference) and fishing mortality F , as well as the corresponding reference points $B_{\text {MSY }}$ and $F_{\text {MSY }}$, the biomass and fishing mortality estimated to lead to the maximum sustainable yield (MSY) (Fig. S1). Only time-series with at least 10 years of data were retained for analysis. We used MSY reference points directly from the assessments, where available, and estimated them using surplus production model fits to stocks' biomass and catch time series (as in refs 10, 11, 32) for the remaining stocks. Of the 253 stocks for which these data were available, 153 stocks, comprising 85 different species, were depleted at some point, and several were depleted multiple times, resulting in a total of 184 depletion events (Supporting Table S1).

## Modeling recovery

Our approach to modeling the recovery process is motivated by the underlying population dynamics of recovering stocks, which can be represented using a stochastic differential equation model for stock biomass dynamics:

$$
\begin{equation*}
d B_{t}=\varphi_{\mathrm{t}} B_{t} d t+\sigma B_{t} d W_{t} \tag{1}
\end{equation*}
$$

where $B_{t}$ is the biomass at time $t$, the function $\varphi_{\mathrm{t}}$, which represents the deterministic component of the biomass dynamics (e.g., survival and reproduction), models the rate of increase of the population, and could, for example take the form of an exponential or logistic growth model with variable fishing mortality (e.g., a Graham-Schaefer surplus production model). $W_{t}$ is a Wiener process with mean 0 and variance coefficient $\sigma$. This continuous-time stochastic process is commonly used to model Brownian motion, and here represents the stochasticity of the biomass
dynamics. Thus, model (1) is a stochastic version of the biomass dynamics models commonly used in fisheries (33), in which the recovery target is approached at a recovery rate determined by $\varphi_{\mathrm{t}}$. Use of this model directly, however, would necessitate specifying the form of $\varphi_{\mathrm{t}}$ and thus determining a priori which (and how) covariates influence the biomass dynamics. This is, however, exactly what we seek to determine empirically, and imposing a particular model for $\varphi_{t}$ may therefore bias our inference. Rather than specifying a specific model for biomass dynamics, our aim is to estimate the importance of covariates that influence these dynamics and the associated stochasticity, and therefore determine the time to recovery. Thus, instead of modeling the time-series themselves, we can model the time to recovery directly by taking the time to recovery as the response variable.

Under the assumption that the a stochastic component of the time-series can be reasonably described by a Wiener process, the time from depletion to recovery follows an inverse-Gaussian (IG) distribution with density

$$
\begin{equation*}
f_{I G}(t)=\frac{c^{\prime}}{\sigma \sqrt{2 \pi}} t^{-3 / 2} \exp \left[-\frac{\left(c^{\prime}-v^{\prime} t\right)^{2}}{2 \sigma^{2} t}\right] \tag{2}
\end{equation*}
$$

where $c^{\prime}=-\log \left(B_{d}\right)$ is a function of the initial biomass at depletion, $B_{d}, v^{\prime}$ is a function of $\varphi_{\mathrm{t}}$, and coefficient $\sigma$ is the Wiener process variance (34). Regardless of the precise model for the biomass, the parameters $c$ ' and $v^{\prime}$ can be interpreted as influencing the 'distance' to the recovery target and the 'recovery rate', respectively. Both parameters are relative to the stochasticity of the biomass dynamics, such that, for instance, the same biomass for highly stochastic dynamics will represent a lower distance from the recovery target since a recovery due to a stochastic event is more likely. $f_{I G}$ thus depends on $\sigma$ through $c^{\prime} / \sigma=c$ and $v^{\prime} / \sigma=v$ (i.e., there are only two free parameters, and we can set $\sigma=1$ without loss of generality (34)), and we may then investigate the importance of covariates in determining the recovery rate and distance from recovery relative to the stochasticity in the time series by placing a regression formulation on each of $c$ and $v(34)$. This means that regression covariates can in theory influence the time to recovery either directly by influencing $c^{\prime}$ or $v^{\prime}$, or by modifying the amount of stochasticity $\sigma$.

The final model for time to recovery for all $i=1 \ldots N$ depletion events can be expressed hierarchically as:

$$
\begin{gathered}
t_{i} \mid v_{i, j}, c_{i} \sim I G\left(v_{i, j}, c_{i}\right) \\
v_{i, j}=\beta^{\prime} X_{i}+\epsilon_{j} \\
c_{i}=-\log \left(B_{d}\right) * \exp \left(\gamma^{\prime} \Xi_{i}\right) \\
\epsilon_{j} \sim N(0, \xi)
\end{gathered}
$$

where, $X_{i}$ and $\Xi_{i}$ are regression covariates influencing the relative recovery rate and the relative distance to recovery, respectively. For the latter, a multiplicative regression formulation assures that the distance remains positive, but is decreased by negative effects of covariates. To circumvent pseudo-replication in stocks with more than one depletion event, we introduce a random effect $\epsilon_{j}$ for the recovery rate of stock j. The probability $\varphi(t)$ that a stock will take at least a time T (e.g., 10 years) to recover is then $F(t)=1-P_{I G}(T \leq t)$, where $P_{I G}$ is the inverse Gaussian cumulative distribution function at T integrated over the random effect (see below).

Importantly, this model form allows for the possibility that some stocks may not recover. Recovery will be increasingly unlikely with an increasingly negative recovery rate and
increasing distance from the recovery target (relative to the stochastic component in the time series). This probability can be directly quantified for our model, and can thus be investigated in terms of regression covariates (see below).

## Survival analysis for stock recoveries

Models of 'time-to-recovery’ often involve censored (i.e., incomplete) data. Specifically, data are censored for each stock that was already depleted at the beginning of its time series (i.e., year of depletion unknown) or was not yet recovered in the final year of its time series (i.e., year of recovery unknown). For such populations, we know the minimum time $t$ of the depletion event, but not the full time of recovery $T$ (which is infinite for populations that never recover). Thus $T>t$ for such depletion events, while $T=t$ for recovered populations in which the full depletion duration is known. These data are most appropriately modeled in a statistical 'time-toevent' framework, which takes into account their incompleteness (35).

To estimate parameters in our model, we need to be able to write the likelihood for all noncensored and censored data. The inverse Gaussian density of time to recovery (2) can be decomposed into the probability of recovering in the time interval $t+\Delta t$ with $\Delta t \rightarrow 0$ given that recovery hasn't taken place up to that point, times the probability that recovery hasn't taken place up to that point. For censored observations, we know only that the depletion event lasted at least time $t$. We thus only have a partial likelihood for these data points. As above, the probability $F_{I G}$ that a stock will take at least a time $t$ to recover is

$$
\mathrm{F}_{\mathrm{IG}}(t)=1-\mathrm{P}_{\mathrm{IG}}(\mathrm{~T} \leq \mathrm{t})=\Phi\left(\frac{\mathrm{c}-v \mathrm{t}}{\sqrt{t}}\right)-\exp (2 c v) \Phi\left(\frac{-\mathrm{c}-v \mathrm{t}}{\sqrt{t}}\right)
$$

The total likelihood is then

$$
L(t)=\prod_{i \in n} \mathrm{~F}_{\mathrm{IG}}(t)^{1-I} f_{I G}(t)^{I}
$$

where $I$ is an indicator taking value 1 for complete and 0 for censored observations.
To make predictions about the statistical population of stocks, we integrate with respect to the random effect, which yields (34)

$$
\mathrm{F}_{\mathrm{U}}(t)=1-\mathrm{P}_{\mathrm{U}}(\mathrm{~T} \leq \mathrm{t})=\Phi\left(\frac{\mathrm{c}-v \mathrm{t}}{\sqrt{t^{2} \xi^{2}+t}}\right)-\exp \left(2 c v+2 c^{2} \xi^{2}\right) \Phi\left(\frac{-\mathrm{c}-2 \mathrm{ct} \xi^{2}-v \mathrm{t}}{\sqrt{t^{2} \xi^{2}+t}}\right)
$$

where the $U$ in the subscript indicates that this quantity is now unconditional, that is, not conditional on the random effect of a particular stock. To investigate if any of our covariates compromise long-term recovery, we take the limit of $F_{U}$ as $t \rightarrow \infty$ (34), which gives

$$
\lim _{t \rightarrow \infty} F_{U}(t)=\Phi\left(\frac{v}{\xi}\right)-\exp \left(2 c v+2 c^{2} \xi^{2}\right) \Phi\left(2 c \xi-\frac{v}{\xi}\right)
$$

Drivers of recovery

We tested a suite of covariates hypothesized to influence recovery. Most obvious is the fishing mortality, F, relative to $\mathrm{F}_{\text {MSY }}$ during the recovery period; here we used the mean $\mathrm{F} / \mathrm{F}_{\text {MSY }}$ during recovery as a potential predictor of the recovery rate. The minimum biomass reached during the depletion may also have an important effect on recovery times, both via density dependent effects on the recovery rate as well as by providing additional information about the distance from the recovery target. High sustained fishing mortality can lead to a number of (potentially hereditary) changes in the demography and physiology of exploited stocks $(18,22)$ that may be directly related to fishing pressure (24) and may limit or enhance the ability of a stock to recover. To account for potential adaptive and evolutionary effects of high and sustained fishing mortality we included the historic fishing intensity as well as the exploitation time as covariates for both the recovery rate and the relative distance. We defined exploitation time as the number of years from the development year, taken as the time from which catches first exceed $20 \%$ of the maximum catch, or the first record in the assessment database if the former was not available, until the year of depletion, and historic fishing intensity as the mean $\mathrm{F} / \mathrm{F}_{\mathrm{MSY}}$ over this period. We hypothesized stocks with long exploitation histories prior to depletion would be more strongly affected by historic fishing intensity, and thus included first order interaction terms of historic fishing intensity with exploitation time (termed 'exploitation history'). All covariates were centered, such that the main effects describe influences at the mean covariate values, and the interaction term describes deviations from mean effects.

We also included the intrinsic rate of increase, $r$, as a covariate to control for 'fast' versus 'slow' life histories in determining the recovery rate (12, 14, 36). The intrinsic rate of increase was estimated for each of these stocks in a taxonomically hierarchical meta-analysis (see Parameters and missing data models below) of surplus production model $r$ estimates (37). We included a binary habitat category for pelagic and demersal species as a predictor of both the recovery rate and relative distance, since the demersal species are often caught in multispecies fishery, where continued bycatch mortality may prolong recoveries (9, 38). Lastly, we included species' trophic level (from Fishbase (39)) in both regressions to account for potentially differential recovery regimes at different levels of the food web. An illustration of covariate distributions is given in Fig. S1. To make effects sizes comparable among continuous and categorical variables, we divided all covariates by twice their standard deviation. We also checked for correlations among all covariates to eliminate the possibility of collinear variables.

Parameter and missing data models
To obtain reasonable estimates of intrinsic rates of increase ( $r$ ), we applied a taxonomic hierarchical meta-analysis method to intrinsic rates of increase $(r)$ estimates obtained from the Schaeffer model. The meta-analysis works by drawing stock level parameters from species level (normal) distributions, which are in turn drawn from family distributions, themselves drawn from an overarching general distribution. The parameter estimates are thus shrunk toward a taxonomically determined mean. This formulation was embedded into the main model in order to let uncertainty about $r$ be directly reflected in parameter estimates (i.e., an 'errors in variables’ approach).

A number of stock assessments included in our analysis did not include a time-series of fishing mortality, and it was therefore impossible to calculate $\mathrm{F} / \mathrm{F}_{\text {MSY }}$ and the historic fishing intensity directly. Our Bayesian computation allowed us to nevertheless include these stocks which may be informative about other parameters, by giving a prior distribution to these parameters. In both cases we used a log-normal distribution as prior, with hyper-priors for its
mean being a normal distribution centered on the log sample mean and with variance of $10^{2}$ and with a vague inverse-gamma hyper-prior (variance $=100$ ) for the variance.

## Priors and implementation

All priors for regression parameters in the final model were vague normal distributions with mean 0 and variance $10^{5}$. For $\xi$ we used a vague Gamma( $0.01,0.01$ ) prior. All covariates were centered and, to compare relative importance of predictors, we divided each covariate by twice its standard deviation. We ran three independent MCMC chains using the WinBUGS package for the computing environment R V.2.14.1 (39), using a burn-in of 50,000 iterations and 500,000 iterations post burn-in. Convergence was assessed visually after thinning the chains, keeping every $100^{\text {th }}$ iteration.

## Model fit

We evaluated the fit of the inverse Gaussian model using standard procedures from survival analysis (41). We calculated Cox-Snell residuals (CSR) as a functional of the joint posterior distribution of model parameters to assess the fit of the Inverse Gaussian distribution to the data. There were slight deviations of the posterior mode from expected CSR values (along the $y=x$ line) for high residuals (Figure S4). Such slight deviations are to be expected at the extremes of CSR values with a relatively limited size of full observations (41). The overall fit seemed reasonable, with the bulk of the posterior distribution of residuals centered along the expected $y=x$ line.

## Robustness to alternative definitions, analyses and potential bias

To inspect whether our results are robust to changes in definitions used for our analysis, we tested a range of alternative models, which gave qualitatively similar results to those discussed above. For example, changing our definitions of depletion to all stocks at biomasses below $\mathrm{B}_{\mathrm{MSY}}$ gave similar, but noisier results, with most of the noise resulting from stocks fluctuating about B/BMSY.

We also explored the effects of generation time on the exploitation history effect by dividing the latter by the former. This analysis gave very similar results to that reported in the main text, as did a model with generation time included as covariate, and a model with recovery time expressed in terms of generation times. The absence of any notable effect from including generation time can be ascribed to its strong correlation with intrinsic rate of increase ( $r^{2}=$ $0.40, p<1.0 e^{-12}$ ). We thus omitted this effect from our overall analysis for reasons of model parsimony.

To ensure the robustness of our results pertaining to exploitation history, we checked for correlations among these variables (i.e., exploitation time and historic fishing intensity), and looked for geographical patterns in their distribution that may hint at these variables being a surrogate for regional differences. We did not find any notable patterns in these analyses (e.g., Fig. S5). We lastly considered an analysis that used the two covariates in question and the stock specific random effect from our analysis reported above to predict deviations from expected biomass increases under a Graham-Schaefer model, as in (12). This analysis confirms the patterns reported in the main text: deviations from expected biomass increases are symmetrically distributed around zero (Fig. S6), with deviations from expectations mainly predicted by exploitation history (Fig. S7); intrinsic rate of increase seems to contribute more than expected (since this parameter is already included in biomass dynamic model used to calculate the
expectation; Fig. S7). This is consistent with adaptations to exploitation happening faster in species with short generation times and high $r$. This analysis also confirms that faster recovery due to exploitation history patterns does not act solely via increased stochasticity in exploited populations (e.g., 42). Since parameters in our model (such as reference points and the intrinsic rate of increase) are estimated directly from such models it may not be too surprising to find that the models fit. Due to the circularity we only introduce this analysis as support for our findings about exploitation time and refrain from further interpretations.

Finally, the stock assessment outputs used as the basis for our analyses are themselves model estimates with potentially large uncertainties. Given the diversity of methods employed to estimate parameters and time-series in stock assessments, these estimation errors likely act to introduce additional uncertainty in our estimates rather than any consistent bias. For some stocks, however, we calculated parameters from the same time-series that were analyzed for recovery, thus leading to a potential upward bias in their relative importance (e.g., intrinsic rate of increase $(r)$ ). There are, however, few alternatives to the calculation of these parameters since stock assessments usually include most of the available information about a species' life-history. We aimed to reduce this potential bias by using a Bayesian meta-analysis technique for $r$ that shrinks values towards a taxonomically determined mean (see Parameter and missing data models above). We further examined the potential bias for MSY reference points calculated from surplus production models, but found little evidence for consistent differences in the assessed depletion levels.

## References and Notes

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Fig. S1
Histograms of covariates used in the analysis. Blue vertical lines indicate mean of covariates over all depletion events.


Fig. $\mathbf{S}^{2}$
Median predicted recovery times (A) and probability of recovering within 10 years (B) as a function of intrinsic rate of increase ( $r$ ) recovering from collapse at no fishing (light green) and fishing at $\mathrm{F}_{\text {MSY }}$ (orange). Areas of overlap between scenarios appear as olive green. For A, relative fishing mortality during the recovery phase is set to zero; all remaining covariates in each plot were fixed at their mean value. Lines are point estimates of median predicted recovery times, shaded regions delimit the $25^{\text {th }}$ to the $75^{\text {th }}$ percentile of predicted recovery times. For B, lines are median estimates; shaded regions are $95 \%$ credible intervals.


Fig. S3
Illustration of definitions used in the analysis. Time series indicate biomass ( B - black circles and solid line) and fishing mortality ( F - open triangles and dotted line) relative to the respective MSY reference points (long dashed line at $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}=1$, short dashed line at $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}=0.5$ ) for Pacific cod (Gadus macrocephalus) off the west coast of Vancouver Island, Canada. The stock is initially depleted in 1984, and then again in 1994. Each recovery period is marked on the biomass time series by cross symbols as opposed to rounds for all other years.


Fig. S4
Residual plot for posterior distribution of Cox-Snell residuals (CSR) from the fitted inverse Gaussian model. The posterior distribution of CSRs (green) was estimated using a normal kernel smoother with $\mathrm{sd}=0.2$. These residuals should scale directly with the estimate of $\mathrm{r}(\mathrm{CSR})$, as indicated by the blue $\mathrm{y}=\mathrm{x}$ line.


Fig. S5
Exploitation history (exploitation time x average relative fishing intensity) prior to depletion among management agencies abbreviated as AFMA: Atlantic Fisheries Management Agency, CCSBT: Commission for the Conservation of Southern Bluefin Tuna, CFP:
Consejo Federal Pesquero, Argentina, DETMCM: Department of Environment and Tourism, Marine and Coastal Management, South Africa, DFO: Department of Fisheries and Oceans, Canada, ICCAT: International Commission for the Conservation of Atlantic Tuna, ICES: International Council for the Exploration of the Seas, MFish: Ministry of Fisheries, New Zealand, NAFO: North Atlantic Fisheries Organization, NMFS: National Marine Fisheries Service, USA.


Fig. S6
Scaled residual plot for observed minus expected biomass increases from a Graham-Schaefer biomass dynamics model. Expected increases were calculated using reported fishing mortality rates.


Fig. S7
Relative importance of predictors of deviations from expected biomass increases from a GrahamSchaefer biomass dynamics model. Positive deviations indicate higher than expected increases in biomass. For a description of predictors, see Fig. 1 of the main text.

## Table S1.

Stock description, country and last assessed year (Year) in the RAM Stock Assessment Database V.1.0 (with addition of ICES status assessments from 2011), as well as binomial species names, and variables used for the recovery analysis. Depletion events are ordered taxonomically (by genus and species). $r$ is the intrinsic growth rate of the stock, $\overline{\boldsymbol{F}_{\boldsymbol{H}}}$ and $\overline{\boldsymbol{F}_{\boldsymbol{D}}}$ are the average relative fishing mortality from the development year to depletion, and during depletion and/or recovery, respectively. Min $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$ is the maximum depletion level, $t$ is the depletion time used for the analysis. Missing cases indicate that this information was not available from the assessments and was imputed (predicted) in the analysis using the posterior predictive distribution for the variable.

| Description | Country | Year | Genus | Species | $r$ | Development Year | $\begin{gathered} \text { Year }<0.5 \\ \mathbf{B}_{\text {MSY }} \end{gathered}$ | $\overline{F_{H}}$ | $\overline{F_{D}}$ | Min B/B $\mathbf{M S Y}$ | t |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sandeel North Sea | Multinational | 2007 | Ammodytes | marinus | 0.89 | 1962 | 2003 | 1.26 | 1.11 | 0.29 | 5 |
| Sablefish Pacific Coast of Canada | Canada | 2004 | Anoplopoma | fimbria | 0.41 | 1962 | 1997 | - | - | 0.27 | 8 |
| Gulf menhaden Gulf of Mexico | USA | 2004 | Brevoortia | patronus | 0.66 | 1952 | 1965 | 7.27 | 5.71 | 0.33 | 4 |
| Atlantic menhaden Atlantic | USA | 2005 | Brevoortia | tyrannus | 0.53 | 1940 | 1963 | - | - | 0.30 | 43 |
| $\begin{gathered} \text { Cusk -FO } \\ 4 \mathrm{X} \end{gathered}$ | Canada | 2007 | Brosme | brosme | 0.47 | 1981 | 2003 | - | - | 0.46 | 5 |
| Black sea bass MidAtlantic Coast | USA | 2007 | Centropristis | striata | 0.63 | 1950 | 1968 | 1.90 | 2.08 | 0.44 | 34 |


| Tanner crab Bering Sea and Aleutian Islands | USA | 2008 | Chionoecetes | bairdi | 0.59 | 1967 | 1978 | 0.64 | 0.36 | 0.10 | 31 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Snow crab <br> Bering Sea | USA | 2008 | Chionoecetes | opilio | 0.17 | 1979 | 1985 | - | - | 0.20 | 3 |
| Snow crab <br> Bering Sea | USA | 2008 | Chionoecetes | opilio | 0.17 | 1979 | 1999 | - | - | 0.36 | 10 |
| New <br> Zealand snapper New Zealand Area 8 | New Zealand | 2005 | Chrysophrys | auratus | 0.08 | 1950 | 1983 | 1.68 | 2.07 | 0.32 | 23 |
| Herring ICES VIa-VIIb-VIIc | Multinational | 2000 | Clupea | harengus | 0.34 | 1950 | 2000 | 2.00 | 2.95 | 0.50 | 1 |
| Herring ICES VIa | Multinational | 2010 | Clupea | harengus | 0.33 | 1950 | 1976 | 2.37 | 2.25 | 0.21 | 34 |
| Herring <br> North Sea | Multinational | 2011 | Clupea | harengus | 0.34 | 1950 | 1996 | 3.85 | 2.14 | 0.46 | 6 |
| Herring ICES 30 | Multinational | 2010 | Clupea | harengus | 0.31 | 1950 | 1979 | 1.11 | 1.04 | 0.36 | 10 |
| Herring ICES 31 | Multinational | 2006 | Clupea | harengus | 0.27 | 1950 | 1994 | 2.89 | 3.41 | 0.24 | 13 |
| Herring Northern Irish Sea | Multinational | 2010 | Clupea | harengus | 0.34 | 1950 | 1961 | 2.84 | 2.68 | 0.13 | 9 |


| Herring Northern Irish Sea | Multinational | 2010 | Clupea | harengus | 0.34 | 1950 | 1978 | 2.68 | 2.31 | 0.13 | 33 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Atlantic herring Northwester n Atlantic Coast | USA | 2005 | Clupea | harengus | 0.36 | 1950 | 1976 | 1.75 | 1.40 | 0.17 | 16 |
| Herring <br> North Sea | Multinational | 2011 | Clupea | harengus | 0.34 | 1950 | 1972 | 5.07 | 3.74 | 0.06 | 15 |
| Herring <br> ICES 25-32 | Multinational | 2010 | Clupea | harengus | 0.27 | 1950 | 1996 | 2.43 | 2.97 | 0.33 | 15 |
| Herring Iceland (Summer spawners) | Multinational | 2010 | Clupea | harengus | 0.33 | 1950 | 1989 | 2.02 | 1.66 | 0.36 | 22 |
| Pacific herring Central Coast | Canada | 2007 | Clupea | pallasii | 0.57 | 1956 | 2006 | 1.15 | 0.34 | 0.30 | 2 |
| Pacific herring West Coast of Vancouver Island | Canada | 2007 | Clupea | pallasii | 0.50 | 1956 | 1995 | 1.00 | 0.41 | 0.03 | 13 |
| Pacific <br> herring <br> Central <br> Coast | Canada | 2007 | Clupea | pallasii | 0.57 | 1956 | 1979 | 1.73 | 0.06 | 0.35 | 2 |
| Pacific herring Straight of Georgia | Canada | 2007 | Clupea | pallasii | 0.54 | 1956 | 1958 | 2.12 | 1.84 | 0.43 | 1 |


| Pacific herring Queen Charlotte Islands | Canada | 2007 | Clupea | pallasii | 0.46 | 1956 | 1986 | 1.63 | 0.42 | 0.08 | 22 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Pacific herring Queen Charlotte Islands | Canada | 2007 | Clupea | pallasii | 0.46 | 1956 | 1951 | 2.10 | 2.58 | 0.14 | 5 |
| Pacific herring West Coast of Vancouver Island | Canada | 2007 | Clupea | pallasii | 0.50 | 1956 | 1966 | 1.78 | 0.86 | 0.27 | 8 |
| Pacific herring Central Coast | Canada | 2007 | Clupea | pallasii | 0.57 | 1956 | 1968 | 2.23 | 0.67 | 0.13 | 7 |
| Pacific herring Straight of Georgia | Canada | 2007 | Clupea | pallasii | 0.54 | 1956 | 1966 | 2.32 | 0.82 | 0.15 | 11 |
| Pacific <br> herring <br> Prince <br> Rupert <br> District | Canada | 2007 | Clupea | pallasii | 0.39 | 1956 | 1953 | 3.95 | 2.31 | 0.31 | 10 |
| Pacific herring Queen Charlotte Islands | Canada | 2007 | Clupea | pallasii | 0.46 | 1956 | 1958 | 3.45 | 1.42 | 0.03 | 23 |
| Pacific herring Prince Rupert District | Canada | 2007 | Clupea | pallasii | 0.39 | 1956 | 1966 | 2.83 | 1.07 | 0.02 | 42 |


| Weakfish Atlantic Coast | USA | 2008 | Cynoscion | regalis | 0.60 | 1950 | 2002 | - | - | 0.13 | 7 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Weakfish Atlantic Coast | USA | 2008 | Cynoscion | regalis | 0.60 | 1950 | 1990 | - | - | 0.39 | 5 |
| Anchovy South Africa | South Africa | 2006 | Engraulis | encrasicolus | 0.43 | 1965 | 1989 | 1.16 | 0.78 | 0.29 | 11 |
| Petrale sole Northern Pacific Coast | USA | 2005 | Eopsetta | jordani | 0.24 | 1942 | 1989 | - | - | 0.45 | 12 |
| Petrale sole Southern Pacific Coast | USA | 2005 | Eopsetta | jordani | 0.19 | 1927 | 1979 | - | - | 0.25 | 26 |
| Snowy grouper Southern Atlantic coast | USA | 2002 | Epinephelus | niveatus | 0.45 | 1975 | 1987 | 2.44 | 3.39 | 0.19 | 16 |
| Pacific cod West Coast of Vancouver Island | Canada | 2002 | Gadus | macrocephalus | 0.43 | 1957 | 1984 | 0.89 | 0.93 | 0.44 | 3 |
| Pacific cod Hecate Strait | Canada | 2005 | Gadus | macrocephalus | 0.42 | 1957 | 2001 | 0.62 | 0.28 | 0.42 | 3 |
| Pacific cod Gulf of Alaska | USA | 2008 | Gadus | macrocephalus | 0.40 | 1965 | 1977 | - | - | 0.47 | 6 |


| Pacific cod <br> West Coast of <br> Vancouver Island | Canada | 2002 | Gadus | macrocephalus | 0.43 | 1957 | 1994 | 1.08 | 0.80 | 0.27 | 8 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Atlantic cod Kattegat | Multinational | 2010 | Gadus | morhua | 0.67 | 1950 | 1982 | 3.02 | 3.24 | 0.04 | 29 |
| Atlantic cod Irish Sea | Multinational | 2010 | Gadus | morhua | 0.65 | 1950 | 1978 | 3.31 | 4.00 | 0.07 | 33 |
| Atlantic cod Georges Bank | USA | 2007 | Gadus | morhua | 0.56 | 1950 | 1984 | 2.51 | 2.62 | 0.07 | 24 |
| Atlantic cod Faroe Plateau | Multinational | 2011 | Gadus | morhua | 0.56 | 1946 | 1990 | 1.83 | 1.47 | 0.29 | 5 |
| Atlantic cod Faroe Plateau | Multinational | 2011 | Gadus | morhua | 0.56 | 1946 | 2004 | 1.87 | 1.81 | 0.30 | 8 |
| Atlantic cod coastal Norway | Multinational | 2010 | Gadus | morhua | 0.32 | 1946 | 2005 | 2.67 | 2.31 | 0.50 | 6 |
| Atlantic cod Baltic Areas 25-32 | Multinational | 2010 | Gadus | morhua | 0.52 | 1950 | 1987 | 3.65 | 3.50 | 0.11 | 24 |
| Atlantic cod West of Scotland | Multinational | 2011 | Gadus | morhua | 0.63 | 1950 | 1990 | 2.65 | 2.66 | 0.11 | 22 |
| Atlantic cod North Sea | Multinational | 2010 | Gadus | morhua | 0.63 | 1950 | 1989 | 2.36 | 2.41 | 0.14 | 22 |


| Atlantic cod -FO 5Zjm | Canada | 2003 | Gadus | morhua | 0.39 | 1950 | 1993 | 3.22 | 2.03 | 0.21 | 11 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Atlantic cod Gulf of Maine | USA | 2007 | Gadus | morhua | 0.52 | 1893 | 1983 | 4.37 | 3.61 | 0.19 | 25 |
| Atlantic cod -FO 3Ps | Canada | 2004 | Gadus | morhua | 0.41 | 1950 | 1991 | 3.20 | 2.20 | 0.16 | 14 |
| Atlantic cod Baltic Areas 25-32 | Multinational | 2010 | Gadus | morhua | 0.52 | 1950 | 1966 | 4.04 | 3.39 | 0.30 | 16 |
| Atlantic cod -FO 3NO | Multinational | 2007 | Gadus | morhua | 0.51 | 1950 | 1960 | 2.12 | 1.47 | 0.02 | 48 |
| Atlantic cod <br> -FO 3Pn4RS | Canada | 2007 | Gadus | morhua | 0.34 | 1950 | 1988 | 4.49 | 3.35 | 0.04 | 20 |
| Atlantic cod Iceland | Multinational | 2011 | Gadus | morhua | 0.50 | 1946 | 1991 | 2.40 | 2.43 | 0.37 | 21 |
| Atlantic cod Baltic Areas 22 and 24 | Multinational | 2011 | Gadus | morhua | 0.75 | 1950 | 1986 | 2.92 | 2.85 | 0.17 | 26 |
| Atlantic cod Northeast Arctic | Multinational | 2010 | Gadus | morhua | 0.57 | 1946 | 1979 | 2.07 | 2.42 | 0.28 | 30 |
| Witch <br> Flounder - <br> FO-5Y | USA | 2007 | Glyptocephalus | cynoglossus | 0.40 | 1960 | 1994 | 2.89 | 3.49 | 0.30 | 14 |


| New <br> Zealand abalone species New Zealand Area PAU 5D | New Zealand | 2006 | Haliotis | iris | 0.27 | 1950 | 2003 | 1.84 | 3.87 | 0.44 | 4 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Flathead sole Bering Sea and Aleutian Islands | USA | 2008 | Hippoglossoides | elassodon | 0.30 | 1954 | 1977 | - | - | 0.14 | 10 |
| American Plaice -FO3LNO | Multinational | 2007 | Hippoglossoides | platessoides | 0.18 | 1957 | 1986 | 3.75 | 4.57 | 0.04 | 22 |
| American Plaice -FO5YZ | USA | 2007 | Hippoglossoides | platessoides | 0.37 | 1957 | 1986 | 2.92 | 2.49 | 0.35 | 22 |
| $\begin{aligned} & \text { Atlantic } \\ & \text { Halibut -FO- } \\ & 5 \mathrm{YZ} \end{aligned}$ | USA | 2007 | Hippoglossus | hippoglossus | 0.40 | 1895 | 1890 | 1.80 | 2.19 | 0.01 | 11 8 |
| American lobster Rhode Island | USA | 2007 | Homarus | americanus | 0.67 | 1950 | 1959 | 1.19 | 1.60 | 0.06 | 49 |
| Orange roughy Southeast Australia | Australia | 2007 | Hoplostethus | atlanticus | 0.08 | 1980 | 1999 | 8.16 | 5.36 | 0.39 | 9 |
| Red rock lobster New Zealand area CRA7 | New Zealand | 2005 | Jasus | edwardsii | 1.12 | 1951 | 1982 | - | - | 0.09 | 24 |
| Red rock lobster New Zealand area CRA5 | New Zealand | 2002 | Jasus | edwardsii | 0.26 | 1947 | 1975 | - | - | 0.12 | 28 |


| Red rock lobster New Zealand area CRA4 | New Zealand | 2005 | Jasus | edwardsii | 0.59 | 1945 | 1977 | - | - | 0.24 | 22 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Red rock lobster New Zealand area CRA2 | New Zealand | 2001 | Jasus | edwardsii | 0.36 | 1945 | 1975 | - | - | 0.19 | 27 |
| Red rock lobster New Zealand area CRA8 | New Zealand | 2005 | Jasus | edwardsii | 1.10 | 1951 | 1987 | - | - | 0.20 | 19 |
| Red rock lobster New Zealand area CRA1 | New Zealand | 2001 | Jasus | edwardsii | 0.34 | 1948 | 1974 | - | - | 0.49 | 22 |
| Rock sole Hecate Strait | Canada | 2001 | Lepidopsetta | bilineata | 0.51 | 1947 | 1945 | 0.31 | 1.50 | 0.33 | 20 |
| Northern rock sole Eastern Bering Sea and Aleutian Islands | USA | 2007 | Lepidopsetta | polyxystra | 0.35 | 1971 | 1975 | 0.84 | 0.61 | 0.16 | 14 |
| Fourspotted megrim ICES VIIIcIXa | Multinational | 2010 | Lepidorhombus | boscii | 0.37 | 1986 | 2000 | 1.92 | 1.66 | 0.47 | 11 |
| Megrim ICES VIIIcIXa | Multinational | 2010 | Lepidorhombus | whiffiagonis | 0.32 | 1950 | 1994 | 1.85 | 1.52 | 0.24 | 17 |
| Yellowfin sole Bering Sea and Aleutian Islands | USA | 2008 | Limanda | aspera | 0.40 | 1977 | 1964 | 3.20 | 2.37 | 0.18 | 15 |


| Yellowtail Flounder FO 3LNO | Multinational | 2009 | Limanda | ferruginea | 0.51 | 1961 | 1976 | 1.44 | 1.18 | 0.20 | 23 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Yellowtail Flounder Southern New EnglandMid Atlantic | USA | 2007 | Limanda | ferruginea | 0.52 | 1939 | 1974 | 5.58 | 5.26 | 0.02 | 34 |
| Yellowtail flounder Cape Cod / Gulf of Maine | USA | 2007 | Limanda | ferruginea | 0.51 | 1941 | 1985 | 6.53 | 6.33 | 0.09 | 23 |
| Yellowtail <br> flounder <br> Georges <br> Bank | USA | 2007 | Limanda | ferruginea | 0.57 | 1948 | 1974 | 4.36 | 4.08 | 0.05 | 34 |
| Tilefish Mid-Atlantic Coast | USA | 2005 | Lopholatilus | chamaeleonticep s | 0.43 | 1950 | 1988 | 1.50 | 1.35 | 0.27 | 18 |
| Mutton snapper Southern Atlantic coast and Gulf of Mexico | USA | 2006 | Lutjanus | analis | 0.44 | 1966 | 1994 | - | - | 0.48 | 11 |
| Red snapper Western Gulf of Mexico | USA | 2003 | Lutjanus | campechanus | 0.38 | 1950 | 1988 | - | - | 0.32 | 16 |
| Red snapper Southern Atlantic coast | USA | 2006 | Lutjanus | campechanus | 0.44 | 1950 | 1968 | 4.80 | 7.53 | 0.02 | 39 |


| Red snapper Eastern Gulf of Mexico | USA | 2003 | Lutjanus | campechanus | 0.39 | 1876 | 1984 | - | - | 0.08 | 20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Hoki Western New Zealand | New Zealand | 2007 | Macruronus | novaezelandiae | 0.49 | 1977 | 2003 | 0.58 | 1.10 | 0.34 | 5 |
| Capelin Iceland | Multinational | 2006 | Mallotus | villosus | 0.66 | 1970 | 2005 | - | - | 0.40 | 2 |
| Capelin <br> Iceland | Multinational | 2006 | Mallotus | villosus | 0.66 | 1970 | 1980 | - | - | 0.42 | 4 |
| Capelin Barents Sea | Multinational | 2006 | Mallotus | villosus | 0.59 | 1970 | 1985 | - | - | 0.02 | 6 |
| Capelin Barents Sea | Multinational | 2006 | Mallotus | villosus | 0.59 | 1970 | 1993 | - | - | 0.04 | 14 |
| Haddock <br> Faroe <br> Plateau | Multinational | 2010 | Melanogrammus | aeglefinus | 0.46 | 1950 | 2008 | 1.84 | 1.26 | 0.31 | 3 |
| Haddock ICES IIIa and North Sea | Multinational | 2010 | Melanogrammus | aeglefinus | 0.51 | 1950 | 2010 | 3.60 | 1.11 | 0.38 | 1 |
| Haddock ICES IIIa and North Sea | Multinational | 2010 | Melanogrammus | aeglefinus | 0.51 | 1950 | 2003 | 3.90 | 1.23 | 0.47 | 2 |
| Haddock ICES IIIa and North Sea | Multinational | 2010 | Melanogrammus | aeglefinus | 0.51 | 1950 | 2007 | 3.71 | 1.33 | 0.36 | 2 |


| Haddock <br> Faroe <br> Plateau | Multinational | 2010 | Melanogrammus | aeglefinus | 0.46 | 1950 | 1991 | 1.89 | 1.22 | 0.29 | 5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Haddock - FO-4X5Y | Canada | 2003 | Melanogrammus | aeglefinus | 0.43 | 1941 | 1973 | 1.49 | 2.32 | 0.45 | 5 |
| Haddock - FO-5Y | USA | 2007 | Melanogrammus | aeglefinus | 0.39 | 1941 | 1986 | 3.34 | 3.97 | 0.09 | 12 |
| Haddock Northeast Arctic | Multinational | 2010 | Melanogrammus | aeglefinus | 0.59 | 1950 | 1977 | 1.66 | 1.22 | 0.16 | 16 |
| Haddock West of Scotland | Multinational | 2006 | Melanogrammus | aeglefinus | 0.54 | 1950 | 1990 | 1.95 | 2.37 | 0.44 | 10 |
| Haddock ICES IIIa and North Sea | Multinational | 2010 | Melanogrammus | aeglefinus | 0.51 | 1950 | 1988 | 4.31 | 4.00 | 0.37 | 11 |
| Haddock Georges Bank | USA | 2007 | Melanogrammus | aeglefinus | 0.49 | 1941 | 1933 | - | - | 0.44 | 28 |
| Haddock Georges Bank | USA | 2007 | Melanogrammus | aeglefinus | 0.49 | 1941 | 1968 | - | - | 0.09 | 38 |
| Haddock - FO-4X5Y | Canada | 2003 | Melanogrammus | aeglefinus | 0.43 | 1941 | 1987 | 2.19 | 1.05 | 0.33 | 17 |
| Haddock Iceland | Multinational | 2011 | Melanogrammus | aeglefinus | 0.55 | 1950 | 1983 | 2.11 | 2.02 | 0.29 | 29 |


| Haddock -FO-5Zejm | Canada | 2003 | Melanogrammus | aeglefinus | 0.41 | 1941 | 1970 | 1.74 | 1.50 | 0.21 | 32 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Whiting ICES VIIe-k | Multinational | 2010 | Merlangius | merlangus | 0.53 | 1951 | 2006 | 3.01 | 2.76 | 0.42 | 5 |
| Whiting ICES VIIe-k | Multinational | 2010 | Merlangius | merlangus | 0.53 | 1951 | 1982 | 3.79 | 3.80 | 0.37 | 11 |
| Argentine hake <br> Northern Argentina | Argentina | 2007 | Merluccius | hubbsi | 0.86 | 1967 | 1988 | 1.97 | 2.03 | 0.12 | 20 |
| Argentine hake Southern Argentina | Argentina | 2008 | Merluccius | hubbsi | 0.93 | 1967 | 1999 | 1.46 | 1.77 | 0.31 | 10 |
| Hake Northeast Atlantic North | Multinational | 2010 | Merluccius | merluccius | 0.49 | 1950 | 1981 | 3.60 | 3.47 | 0.19 | 30 |
| Deep-water cape hake South Africa | South Africa | 2008 | Merluccius | paradoxus | 0.61 | 1952 | 1981 | - | - | 0.48 | 15 |
| Whiting Northeast Atlantic | Multinational | 2007 | Micromesistius | poutassou | 0.38 | 1979 | 2007 | 2.26 | - | 0.41 | 1 |
| Whiting Northeast Atlantic | Multinational | 2007 | Micromesistius | poutassou | 0.38 | 1979 | 1981 | 2.06 | 2.13 | 0.25 | 21 |
| Atlantic croaker MidAtlantic Coast | USA | 2002 | Micropogonias | undulatus | 0.51 | 1954 | 1973 | 0.74 | 0.54 | 0.21 | 14 |


| Gag Gulf of Mexico | USA | 2004 | Mycteroperca | microlepis | 0.31 | 1975 | 1979 | 1.22 | 1.76 | 0.44 | 26 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Jackass morwong Southeast Australia | Australia | 2007 | Nemadactylus | macropterus | 0.12 | 1946 | 1997 | 3.93 | 4.26 | 0.26 | 11 |
| Tiger flathead Southeast Australia | Australia | 2006 | Neoplatycephalu <br> s | richardsoni | 0.24 | 1919 | 1949 | 1.24 | 0.25 | 0.39 | 31 |
| Red porgy Southern Atlantic coast | USA | 2004 | Pagrus | pagrus | 0.16 | 1972 | 1990 | 1.93 | 2.17 | 0.26 | 15 |
| Northern shrimp Gulf of Maine | USA | 2009 | Pandalus | borealis | 0.61 | 1968 | 1998 | 0.50 | 0.38 | 0.26 | 8 |
| Summer flounder Mid-Atlantic Coast | USA | 2007 | Paralichthys | dentatus | 0.44 | 1950 | 1982 | - | - | 0.12 | 26 |
| Red king crab Pribilof Islands | USA | 2009 | Paralithodes | camtschaticus | 0.72 | 1950 | 1983 | - | - | 0.01 | 11 |
| Red king crab Bristol Bay | USA | 2008 | Paralithodes | camtschaticus | 0.22 | 1950 | 1985 | 3.98 | 2.51 | 0.39 | 17 |
| Blue king crab Pribilof Islands | USA | 2008 | Paralithodes | platypus | 0.60 | 1975 | 2000 | - | - | 0.03 | 9 |
| Blue king crab Saint Matthews Island | USA | 2008 | Paralithodes | platypus | 0.59 | 1975 | 1985 | - | - | 0.37 | 5 |


| Blue king crab Pribilof Islands | USA | 2008 | Paralithodes | platypus | 0.60 | 1975 | 1986 | - | - | 0.15 | 10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Blue king crab Saint Matthews Island | USA | 2008 | Paralithodes | platypus | 0.59 | 1975 | 1999 | - | - | 0.37 | 8 |
| English sole <br> Hecate Strait | Canada | 2001 | Parophrys | vetulus | 0.64 | 1948 | 1961 | 1.44 | 1.32 | 0.40 | 6 |
| Sea scallop Georges Bank | USA | 2006 | Placopecten | magellanicus | 0.55 | 1953 | 1982 | 3.46 | 2.58 | 0.12 | 19 |
| Sea scallop Mid-Atlantic Coast | USA | 2006 | Placopecten | magellanicus | 0.58 | 1953 | 1975 | 1.14 | 2.13 | 0.09 | 31 |
| Starry flounder Southern Pacific Coast | USA | 2005 | Platichthys | stellatus | 0.36 | 1970 | 1980 | - | - | 0.34 | 5 |
| European Plaice ICES VIIe | Multinational | 2010 | Pleuronectes | platessa | 0.54 | 1950 | 2009 | 2.37 | 1.74 | 0.48 | 1 |
| European Plaice ICES VIIf-g | Multinational | 2006 | Pleuronectes | platessa | 0.65 | 1950 | 2002 | 1.84 | 1.89 | 0.42 | 5 |
| European <br> Plaice Irish Sea | Multinational | 2006 | Pleuronectes | platessa | 0.59 | 1950 | 1995 | 1.37 | 0.64 | 0.45 | 11 |
| European <br> Plaice Irish Sea | Multinational | 2006 | Pleuronectes | platessa | 0.59 | 1950 | 1976 | - | - | 0.39 | 12 |


| $\begin{aligned} & \text { Pollock -FO- } \\ & 4 \mathrm{X} 5 \mathrm{YZ} \end{aligned}$ | Canada | 2006 | Pollachius | virens | 0.47 | 1980 | 1994 | 2.19 | 2.07 | 0.18 | 13 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Pollock Northeast Arctic | Multinational | 2010 | Pollachius | virens | 0.52 | 1950 | 1984 | 1.53 | 1.73 | 0.33 | 15 |
| Pollock ICES IIIa, VI and North Sea | Multinational | 2010 | Pollachius | virens | 0.44 | 1950 | 1967 | 2.28 | 2.03 | 0.26 | 44 |
| Pollock <br> Faroe <br> Plateau | Multinational | 2010 | Pollachius | virens | 0.40 | 1950 | 1961 | 0.61 | 1.56 | 0.44 | 26 |
| Winter <br> Flounder <br> Southern New <br> EnglandMid Atlantic | USA | 2007 | Pseudopleuronectes | americanus | 0.53 | 1950 | 1981 | 3.66 | 3.58 | 0.05 | 27 |
| Winter <br> Flounder - <br> FO-5Z | USA | 2006 | Pseudopleuronectes | americanus | 0.58 | 1950 | 1985 | 2.44 | 2.03 | 0.21 | 22 |
| Winter flounder Rhode Island | USA | 2007 | Pseudopleuronectes | americanus | 0.49 | 1950 | 1971 | 1.83 | 1.87 | 0.15 | 37 |
| Greenland halibut -FO 23KLMNO | Multinational | 2006 | Reinhardtius | hippoglossoides | 0.34 | 1964 | 1995 | 2.02 | 2.57 | 0.38 | 12 |
| Greenland halibut Northeast Arctic | Multinational | 2010 | Reinhardtius | hippoglossoides | 0.27 | 1964 | 1975 | 2.25 | 1.92 | 0.15 | 36 |
| Arrowtooth flounder Gulf of Alaska | USA | 2010 | Reinhardtius | stomias | 0.11 | 1955 | 1961 | - | - | 0.46 | 20 |


| common <br> gemfish <br> Southeast <br> Australia | Australia | 2007 | Rexea | solandri | 0.14 | 1974 | 1991 | 0.62 | 0.35 | 0.13 | 17 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sardine South Africa | South Africa | 2006 | Sardinops | sagax | 0.52 | 1977 | 1984 | - | 0.56 | 0.06 | 13 |
| Pacific chub mackerel Pacific Coast | USA | 2008 | Scomber | japonicus | 0.40 | 1929 | 2000 | - | - | 0.22 | 9 |
| Pacific chub mackerel Pacific Coast | USA | 2008 | Scomber | japonicus | 0.40 | 1929 | 1948 | - | - | 0.09 | 29 |
| Atlantic mackerel Gulf of Maine / Cape Hatteras | USA | 2004 | Scomber | scombrus | 0.39 | 1965 | 1962 | 0.27 | 0.45 | 0.46 | 6 |
| Spanish mackerel Southern Atlantic Coast | USA | 2007 | Scomberomorus | maculatus | 0.38 | 1951 | 1961 | 2.36 | 2.33 | 0.15 | 47 |
| Cabezon <br> Southern <br> California | USA | 2005 | Scorpaenichthys | marmoratus | 0.28 | 1972 | 1983 | 1.18 | 1.86 | 0.13 | 23 |
| Pacific ocean perch Pacific Coast | USA | 2007 | Sebastes | alutus | 0.07 | 1962 | 1994 | 1.76 | 0.64 | 0.46 | 14 |
| Pacific Ocean perch Eastern Bering Sea and Aleutian Islands | USA | 2009 | Sebastes | alutus | 0.24 | 1963 | 1977 | - | - | 0.22 | 19 |


| Pacific ocean perch Gulf of Alaska | USA | 2010 | Sebastes | alutus | 0.14 | 1963 | 1977 | - | - | 0.23 | 21 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Darkblotche d rockfish Pacific Coast | USA | 2007 | Sebastes | crameri | 0.06 | 1951 | 1994 | 1.62 | 1.68 | 0.35 | 14 |
| Acadian redfish Gulf of Maine / Georges Bank | USA | 2007 | Sebastes | fasciatus | 0.20 | 1936 | 1957 | - | - | 0.03 | 51 |
| Cowcod Southern California | USA | 2007 | Sebastes | levis | 0.22 | 1916 | 1981 | 2.36 | 6.19 | 0.04 | 27 |
| Blue rockfish California | USA | 2007 | Sebastes | mystinus | 0.20 | 1951 | 1985 | 1.48 | 1.83 | 0.25 | 23 |
| Bocaccio Southern Pacific Coast | USA | 2006 | Sebastes | paucispinis | 0.23 | 1951 | 1960 | 2.16 | 1.67 | 0.47 | 5 |
| Bocaccio Southern Pacific Coast | USA | 2006 | Sebastes | paucispinis | 0.23 | 1951 | 1984 | 2.52 | 2.25 | 0.14 | 23 |
| Canary rockfish Pacific Coast | USA | 2009 | Sebastes | pinniger | 0.18 | 1944 | 1992 | 0.68 | 0.74 | 0.32 | 18 |
| Yelloweye rockfish Pacific Coast | USA | 2006 | Sebastes | ruberrimus | 0.07 | 1951 | 1994 | 3.55 | 4.16 | 0.31 | 13 |
| Redfish species -FO 3LN | Multinational | 2008 | Sebastes | spp | 0.35 | 1959 | 1992 | 1.06 | 0.67 | 0.35 | 8 |


| Dusky rockfish Gulf of Alaska | USA | 2007 | Sebastes | variabilis | 0.15 | 1951 | 1978 | 0.34 | 0.47 | 0.46 | 19 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Blue Warehou Eastern half of Southeast Australia | Australia | 2006 | Seriolella | brama | 0.38 | 1987 | 2001 | 0.70 | 0.55 | 0.42 | 6 |
| Blue Warehou Western half of Southeast Australia | Australia | 2006 | Seriolella | brama | 0.42 | 1987 | 1994 | 1.43 | 2.13 | 0.21 | 13 |
| School whiting Southeast Australia | Australia | 2007 | Sillago | flindersi | 0.48 | 1974 | 2000 | 0.90 | 0.93 | 0.49 | 8 |
| common European sole Irish Sea | Multinational | 2011 | Solea | vulgaris | 0.43 | 1950 | 2004 | 1.97 | 1.53 | 0.24 | 8 |
| common <br> European sole ICES <br> Kattegat and Skagerrak | Multinational | 2010 | Solea | vulgaris | 0.49 | 1950 | 1984 | 1.38 | 1.38 | 0.33 | 6 |
| common European sole Western English Channel | Multinational | 2010 | Solea | vulgaris | 0.32 | 1950 | 1969 | 0.85 | 1.74 | 0.48 | 42 |
| Sprat ICES <br> Baltic Areas 22-32 | Multinational | 2010 | Sprattus | sprattus | 0.44 | 1966 | 1979 | 2.04 | 1.36 | 0.33 | 4 |
| Sprat ICES Baltic Areas 22-32 | Multinational | 2010 | Sprattus | sprattus | 0.44 | 1966 | 1986 | 1.67 | 1.41 | 0.35 | 6 |


| Tautog Rhode Island | USA | 2007 | Tautoga | onitis | 0.54 | 1976 | 1992 | 1.61 | 1.20 | 0.34 | 16 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Walleye pollock Eastern Bering Sea | USA | 2008 | Theragra | chalcogramma | 0.53 | 1967 | 1964 | 0.38 | 1.00 | 0.23 | 18 |
| Southern bluefin tuna Southern Oceans | Multinational | 2009 | Thunnus | maccoyii | 0.37 | 1978 | 1990 | - | - | 0.17 | 20 |
| Bluefin tuna Eastern Atlantic | Multinational | 2007 | Thunnus | thynnus | 0.40 | 1950 | 2004 | 3.47 | 5.91 | 0.34 | 4 |
| Bluefin tuna Western Atlantic | Multinational | 2007 | Thunnus | thynnus | 0.35 | 1950 | 1989 | 2.99 | 1.68 | 0.43 | 19 |
| Norway pout North Sea | Multinational | 2010 | Trisopterus | esmarkii | 0.50 | 1968 | 2003 | 3.27 | 1.80 | 0.32 | 5 |
| White hake Georges Bank / Gulf of Maine | USA | 2007 | Urophycis | tenuis | 0.29 | 1950 | 1994 | 1.51 | 1.86 | 0.24 | 14 |
| White hake Georges Bank / Gulf of Maine | USA | 2007 | Urophycis | tenuis | 0.29 | 1950 | 1963 | 1.30 | 0.81 | 0.36 | 12 |

## Table S2.

Quantiles of regression parameter posterior distributions for standardized predictors of recovery schedules. The table is divided into predictors of the recovery rate, the relative distance from recovery target, and the lower level regression for Minimum $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$. Median point estiamtes appear in bold.

| Parameter | Predictor | 2.5\% | Estimate <br> Median | $97.5 \%$ | $P(\beta>0)$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | F/F $\mathrm{F}_{\text {MSY }}$ | -0.40 | $\mathbf{- 0 . 2 4}$ | -0.11 | 0.00 |
|  | Exploitation History | -0.35 | $\mathbf{- 0 . 2 1}$ | 0.07 | 0.07 |
| recovery rate | Exploitation Time | -0.19 | $\mathbf{0 . 1 5}$ | 0.30 | 0.84 |
| $(v)$ | Intrinsic Rate of Increase | 0.04 | $\mathbf{0 . 1 4}$ | 0.28 | 1.00 |
|  | Historic Fishing Intensity | -0.08 | $\mathbf{0 . 0 7}$ | 0.23 | 0.83 |
|  | Trophic Level | -0.17 | $\mathbf{- 0 . 0 5}$ | 0.06 | 0.21 |
|  | Habitat | -0.08 | $\mathbf{0 . 0 2}$ | 0.14 | 0.66 |
|  | Minimum B/B MSY $^{2}$ | -0.08 | $\mathbf{0 . 0 2}$ | 0.12 | 0.65 |
| relative | Minimum B/B B $_{\text {MSY }}$ | -0.59 | $\mathbf{- 0 . 4 3}$ | -0.26 | 0.00 |
| distance | Historic Fishing Intensity | -0.33 | $\mathbf{- 0 . 1 9}$ | -0.03 | 0.01 |
| from | Habitat | -0.23 | $\mathbf{- 0 . 0 8}$ | 0.10 | 0.21 |
| recovery | Exploitation Time | -0.22 | $\mathbf{- 0 . 0 5}$ | 0.15 | 0.34 |
| target (c) | Trophic Level | -0.13 | $\mathbf{0 . 0 4}$ | 0.23 | 0.63 |
|  | Exploitation History | -0.24 | $\mathbf{0 . 0 0}$ | 0.28 | 0.50 |

